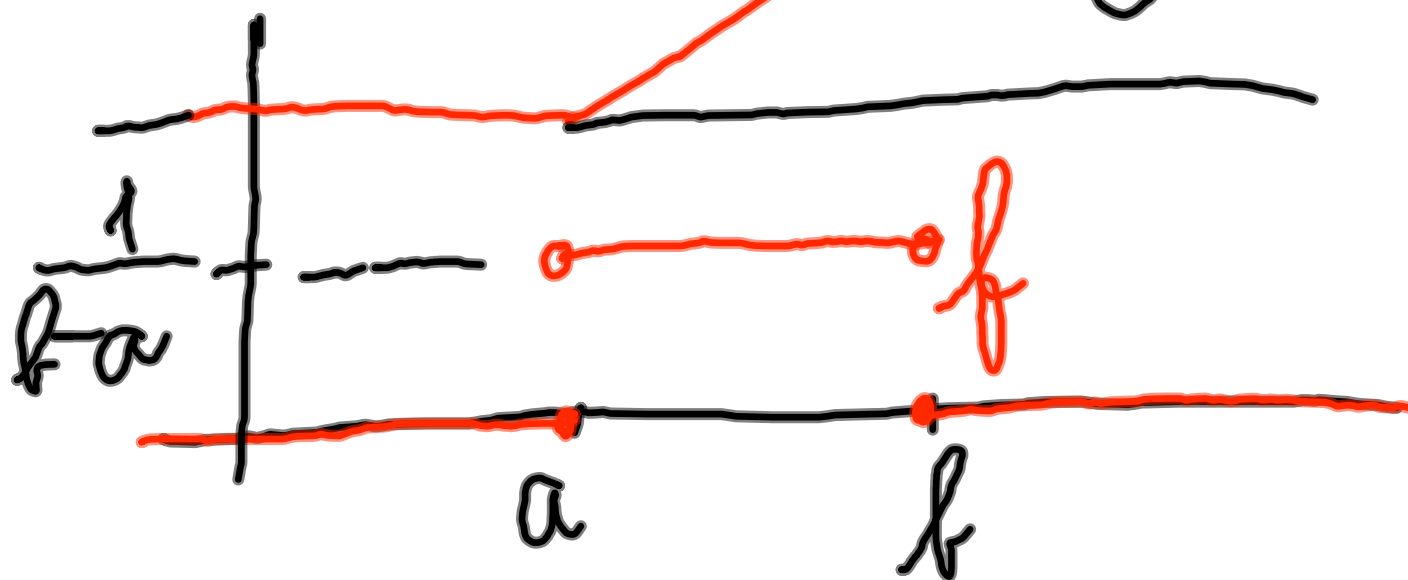


$$P(a \leq X < b) = \int_a^b f(x) dx$$
$$P(a < X \leq b)$$
$$\vdots$$

Def.  $X$  má rovnorné  
rozdělení na  $(a, b)$

$$f(x) = \begin{cases} 0 & x \leq a \\ \frac{1}{b-a} & x \in (a, b) \\ 0 & x \geq b \end{cases}$$



$$\int_{-\infty}^{\infty} f(x) dx = 1$$

Def.  $X$  má normální  
rozdělení s parametry

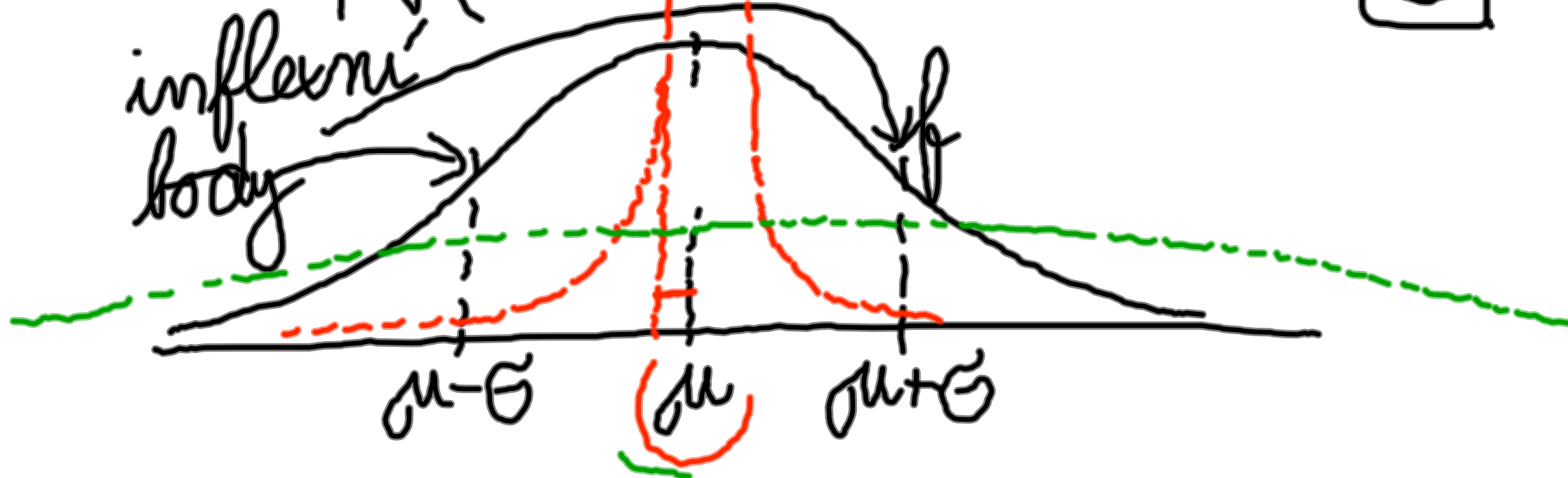
$$\mu \in \mathbb{R}, \sigma \in \mathbb{R}^+$$

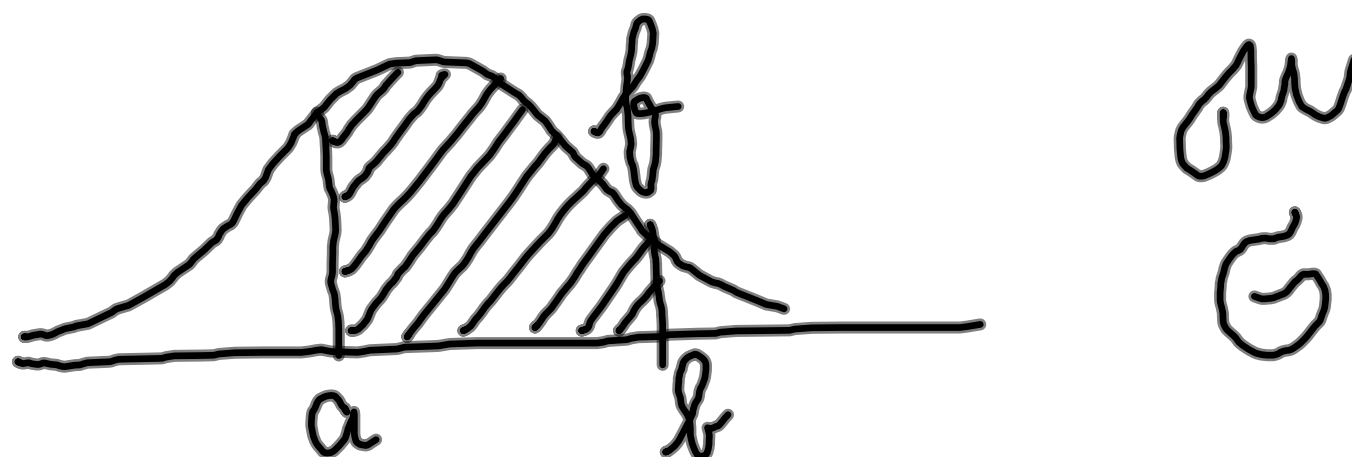
pokud pro hustotu platí

$$f(x) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$\int_{-\infty}^{\infty} f(x) dx = 1$$

$X \sim N(\mu, \sigma)$        $N(5, 3)$   
 $N(\mu, \sigma^2)$





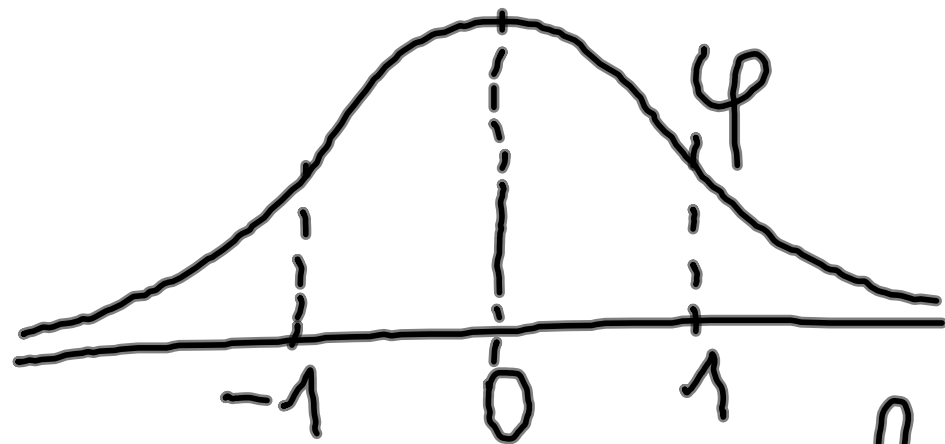
$$\mu = 0$$

$$\sigma = 1$$

normované normální

$$N(0,1)$$

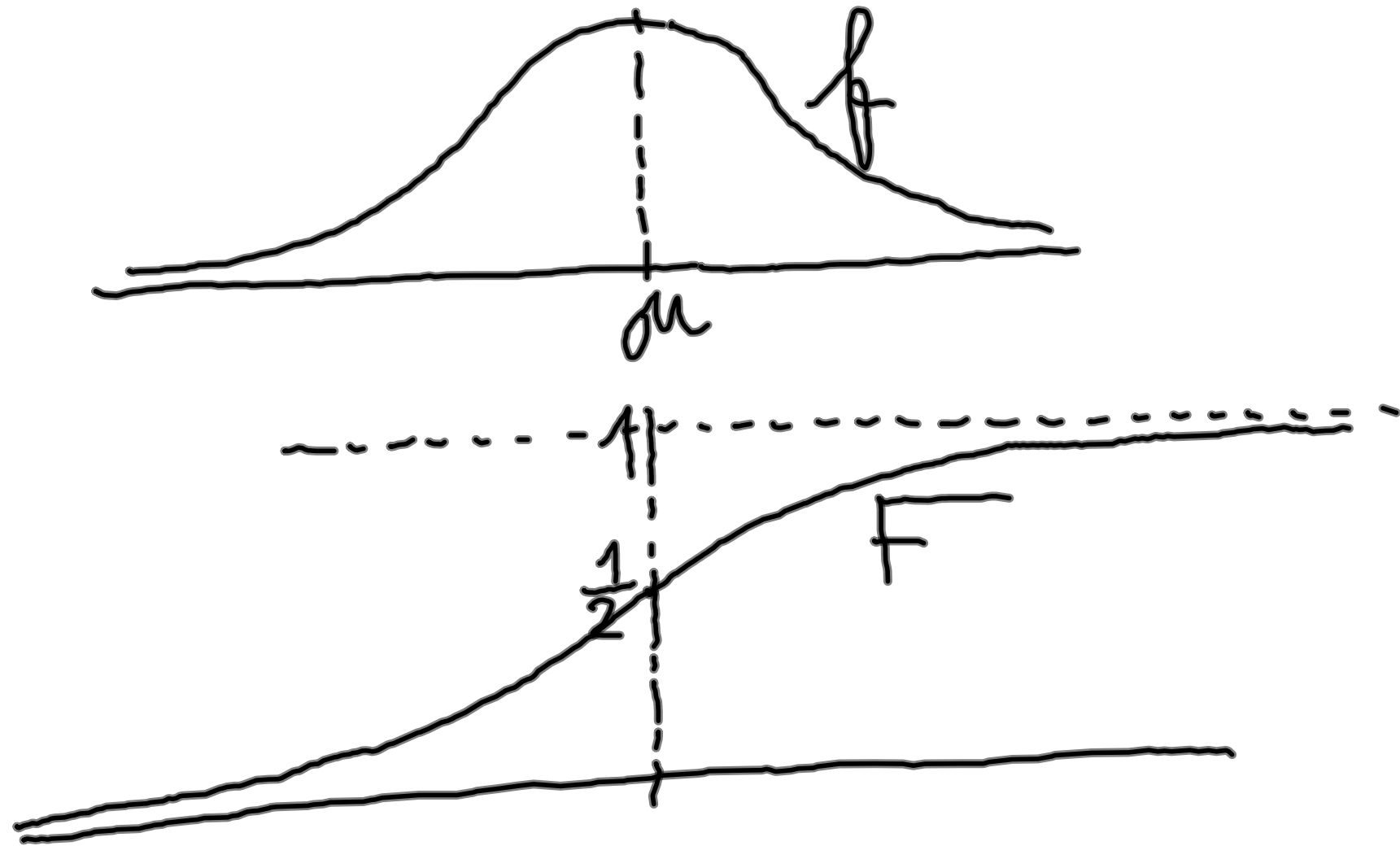
hustota  $\varphi(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$



distribuční funkce

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{u^2}{2}} du$$

tabulované



Trázení:  $X \sim N(\mu, \sigma)$ , F... d.f.

Platí:

$$F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$$

Důkaz:

$$F(x) = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$\frac{x-\mu}{\sigma} = u$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\frac{x-\mu}{\sigma}} e^{-\frac{u^2}{2}} du =$$





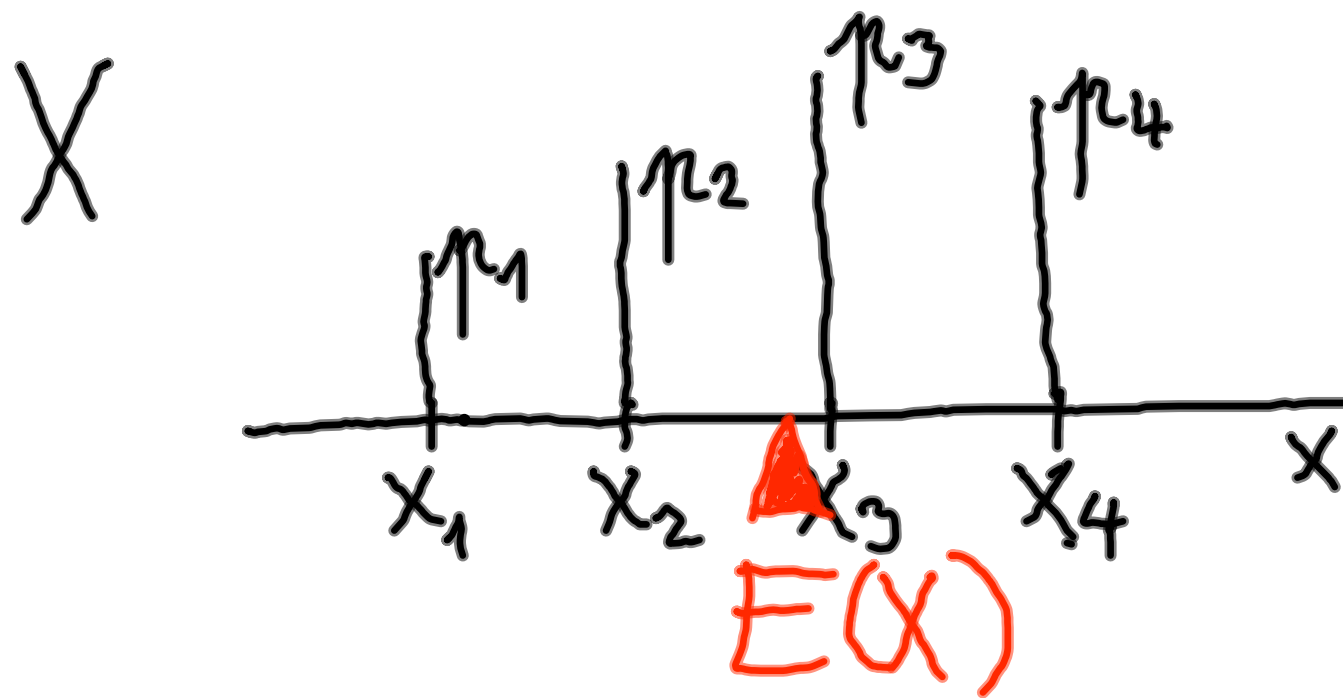
# Číselné charakteristiky

Průměrná hodnota  $E(X)$

$X$  diskrétní

$$P(X = x_i) = p_i, \quad i = 1, 2, \dots$$

$$E(X) = \sum_i x_i p_i$$



Charakteristika  
řolohy  
 $E(Y) = E(X) + a$

Prům. 2

$$y_i = x_i + a$$

$X$

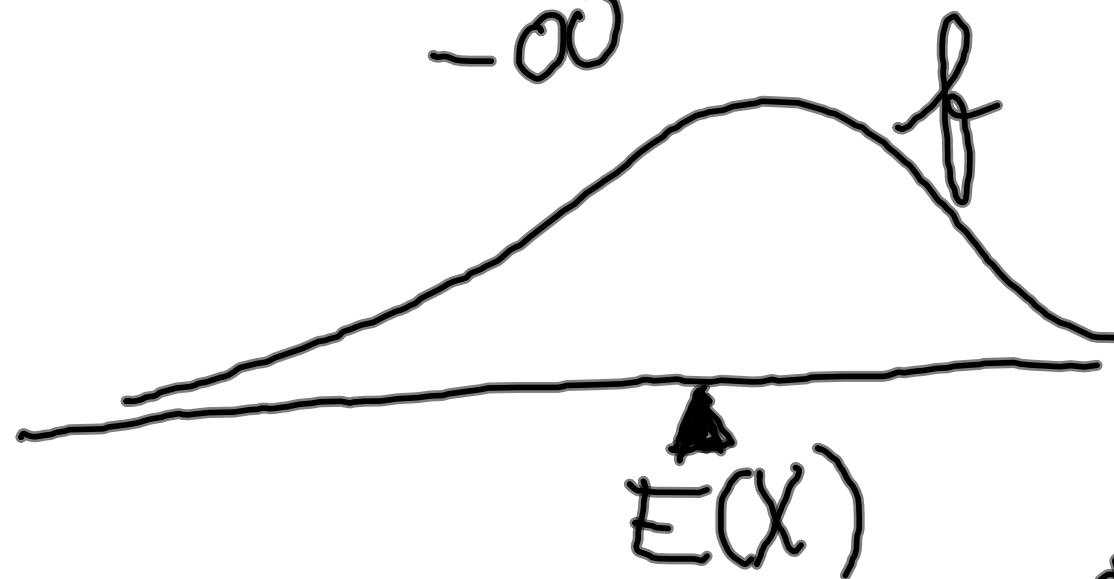
$$Y = X + a$$

$$P(Y = y_i) = P(X = x_i) = p_i$$

$$E(Y) = \sum_{i=1}^m y_i p_i = \sum_{i=1}^m (x_i + a) p_i = \underbrace{\sum_{i=1}^m x_i p_i}_{E(X)} + a \underbrace{\sum_{i=1}^m p_i}_1$$

$X$  je spojitel s hustotou  $f$

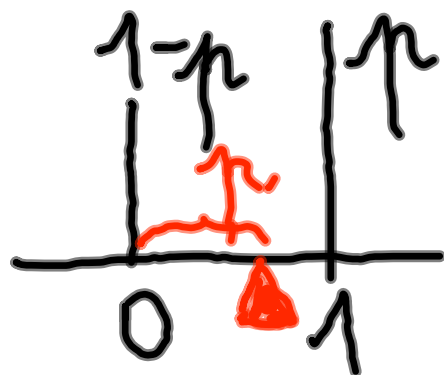
$$E(X) = \int_{-\infty}^{\infty} x f(x) dx$$



$$X+a$$
$$E(X+a) = E(X) + a$$

Př.  $X$  má alternativní rozdělení

$$E(X) = ?$$



$$E(X) = 0 \cdot (1-p) + 1 \cdot p = p$$

Pr.  $X \sim \text{Bi}(n, p)$   $P(A) = p$

$$k = 0, 1, 2, \dots, n$$

$$P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$$

$$E(X) = ?$$

$$E(X) = \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k}$$

$$\begin{aligned} k \binom{n}{k} &= k \frac{n!}{k! (n-k)!} = \frac{n}{(k-1)! (n-k)!} \\ &= n \binom{n-1}{k-1} \end{aligned}$$

$$\begin{aligned}
 E(X) &= \sum_{k=0}^m k \binom{m}{k} p^k (1-p)^{m-k} = \\
 &= m \sum_{k=1}^m \binom{m-1}{k-1} p^k (1-p)^{m-k} = \underline{m} p \\
 &= m \sum_{k=0}^{m-1} \binom{m-1}{k} p^{k+1} (1-p)^{m-k-1} \\
 &= m p \sum_{k=0}^{m-1} \binom{m-1}{k} p^k (1-p)^{(m-1)-k} = m p \cdot 1 = \underline{m p}
 \end{aligned}$$

The derivation shows the simplification of the binomial distribution's mean. Key steps include:

- Using the identity  $k \binom{m}{k} = m \binom{m-1}{k-1}$  (highlighted in red).
- Shifting the summation index from  $k=1$  to  $k=0$ .
- Factoring out  $p$  and recognizing the sum as the binomial expansion of  $(p + (1-p))^{m-1} = 1$  (highlighted in red).



Pi.

$$X \sim N(\mu, \sigma)$$

$$E(X) = \mu$$

